Overview

Robert A. Miller

Structural Econometrics

October 2023

-

Image: A math a math

Lectures on Structural Econometrics

Website, topics and themes

- The lecture material, some assignments and background reading for these 28 sessions can be found at:
 - http://comlabgames.com/structuraleconometrics/
- There are two sets of lectures with four segments in the first group:
 - Introduction to Structural Econometrics
 - ② Summarizing the Data
 - Probability
 - Asymptotic Theory for Nonlinear Models
- There are three segments in the second set of lectures.
 - Dynamic Discrete Choice
 - 2 Market Microstructure
 - Optimal Contracting
- Throughout these lectures we will imagine the data is generated by a model, and embrace the classical laws of probability and statistics.

General approach to estimation and testing

• For the most part we assume the model comes from economics:

- Individuals solve dynamic optimization problems.
- Groups of individuals or firms play a noncooperative game using equilibrium strategies.
- Asymmetrically informed individuals optimally contract with each other.
- Individuals and firms make consumption and production choices in competitive equilibrium.
- To help understand how economic models provide the basis for estimation and testing we introduce the course by analyzing some of the first structural econometric models in:
 - dynamic discrete choice
 - competitive equilibrium models with continuous choices
 - market microstructure
 - optimal contracting with moral hazard.

Introduction to Structural Econometrics Modeling

Data generating process

- The data typically comprise a sample of individuals for which there are records on some of their:
 - background characteristics
 - choices
 - outcomes from those choices.
- What are the challenges to making predictions and testing hypotheses when we take this approach?
 - The choices and outcomes of economic models are typically nonlinear in the underlying parameters of the model we wish to estimate.
 - The data variables on background, choices and outcomes might be an incomplete description about what is relevant to the model.

Dynamic Discrete Choice

- Each period t ∈ {1, 2, ..., T} for T ≤ ∞, an individual chooses among J mutually exclusive actions.
- Let d_{jt} equal one if action j ∈ {1,..., J} is taken at time t and zero otherwise:

$$d_{jt} \in \{0,1\}$$
 $\sum_{j=1}^J d_{jt} = 1$

- At an abstract level assuming that choices are mutually exclusive is innocuous, because two combinations of choices sharing some features but not others can be interpreted as two different choices.
- For example in a female labor supply and fertility model, suppose:
 - $j \in \{(work, no birth), (work, birth), (no work, no birth), (no work, birth)\}$

Image: Image:

Dynamic Discrete Choice

Information and states

- Suppose that actions taken at time t can potentially depend on the state z_t ∈ Z.
- For Z finite denote by $f_{jt}(z_{t+1}|z_t)$, the probability of z_{t+1} occurring in period t + 1 when action j is taken at time t.
- For example in the example above, suppose $z_t = (w_t, k_t)$ where:
 - $k_t \in \{0, 1, \ldots\}$ are the number of births before t
 - $w_t \equiv d_{1,t-1} + d_{2,t-1}$, so $w_t = 1$ if the female worked in period t 1, and $w_t = 0$ otherwise.
- Note that Z must be defined compatible to the transition matrix: for example setting z_t = (w_t, k_t) where k_t ∈ {0, 1, ...} are the number of births before t − 1, is incompatible with assumption about transitions and choices.
- With up to 5 offspring, 3 levels of experience, the number of states including age (say 50 years) is 750. Add in 4 levels of education (less than high school, high school, some college and college graduate) and 3 racial categories, increases this number to 9000.

Miller (Structural Econometrics)

Dynamic Discrete Choice

Large but sparse matrices

- When Z is finite there is a $Z \times Z$ transition matrix for each (j, t).
- In many applications the matrices are sparse.
- In the example above they have $9,000^2 = 81$ million cells.
- However households can only increase the number of kids one at time.
- They can only increase or decrease their work experience by one unit at most.
- Hence there are at most six cells they can move from (w_t, k_t) :

$$\left\{\begin{array}{l} (w_t, k_t), (w_t, k_t+1), (w_t+1, k_t), \\ (w_t+1, k_t+1), (w_t-1, k_t), (w_t-1, k_t+1) \end{array}\right\}$$

- Therefore a transition matrix has at most 54,000 nonzero elements, and all the nonzero elements are one.
- Given a deterministic sequence of actions sequentially taken over S periods, we can form the S period transition matrix by producting the one period transitions.

More on information and states

- If Z is a Euclidean space $f_{jt}(z_{t+1}|z_t)$ is the probability (density function) of z_{t+1} occurring in period t + 1 when j is picked at time t.
- With almost identical notation we could model z_t ∈ Z_t and in this way generalize from states of the world to histories, or information known at t, or t-measurable events.
- For example in a health application we might define z_t ≡ {h_s}^{t-1}_{s=1} as a medical record with h_s ∈ {healthy at s, sick at s}.

Dynamic Discrete Choice Models

Preferences and expected utility

- The individual's current period payoff from choosing *j* at time *t* is determined by *z*_t, which is revealed to the individual at the beginning of the period *t*.
- The current period payoff at time t from taking action j is $u_{jt}(z_t)$.
- Given choices (d_{1t}, \ldots, d_{Jt}) in each period $t \in \{1, 2, \ldots, T\}$ and each state $z_t \in Z$ the individual's expected utility is:

$$E\left\{\sum_{t=1}^{T}\sum_{j=1}^{J}\beta^{t-1}d_{jt}u_{jt}(z_{t})|z_{1}\right\}$$

where $\beta \in (0, 1)$ is the subjective discount factor, and at each period t the expectation is taken over z_2, \ldots, z_T .

• Formally β is redundant if u is subscripted by t; we typically include a geometric discount factor so that infinite sums of utility are bounded, and the optimization problem is well posed.

Dynamic Discrete Choice Models

Value Function

- Write the optimal decision at period t as a decision rule denoted by $d_t^o(z_t)$ formed from its elements $d_{jt}^o(z_t)$.
- Let $V_t(z_t)$ denote the value function in period t, conditional on behaving according to the optimal decision rule:

$$W_t(z_t) \equiv E\left[\sum_{\tau=t}^T \sum_{j=1}^J \beta^{\tau-t} d_{j\tau}^o(z_\tau) u_{j\tau}(z_\tau) | z_t\right]$$

• In terms of period t + 1:

$$\beta V_{t+1}(z_{t+1}) \equiv \beta E \left\{ \sum_{\tau=t+1}^{T} \sum_{j=1}^{J} \beta^{\tau-t-1} d_{j\tau}^{o}(z_{\tau}) u_{j\tau}(z_{\tau}) | z_{t+1} \right\}$$

• Appealing to Bellman's (1958) principle we obtain, when Z is finite:

$$\begin{aligned} V_t(z_t) &= \sum_{j=1}^J d_{jt}^o u_{jt}(z_t) \\ &+ \sum_{j=1}^J d_{jt}^o \sum_{z \in Z} E\left[\sum_{\tau=t+1}^T \sum_{j=1}^J \beta^{\tau-t} d_{j\tau}^o(z_\tau) \, u_{j\tau}(z_\tau) \, |z\right] f_{jt}(z|z_t) \\ &= \sum_{j=1}^J d_{jt}^o \left[u_{jt}(z_t) + \beta \sum_{z \in Z} V_{t+1}(z) f_{jt}(z|z_t) \right] \end{aligned}$$

• A similar expression holds when Z is Euclidean using an integral.

Dynamic Discrete Choice Models

- To compute the optimum for T finite, we first solve a static problem in the last period to obtain d^o_T(z_T) for all z_T ∈ Z.
- Applying backwards induction $i \in \{1, ..., J\}$ is chosen to maximize:

$$u_{it}(z_{t}) + E\left\{\sum_{\tau=t+1}^{T}\sum_{j=1}^{J}\beta^{\tau-t-1}d_{j\tau}^{o}\left(z_{\tau}\right)u_{j\tau}(z_{\tau})\left|z_{t},d_{it}=1\right.\right\}$$

- In the stationary infinite horizon case we assume $u_{jt}(z) \equiv u_j(z)$ and that $u_j(z) < \infty$ for all (j, z).
- Consequently expected utility each period is bounded and the contraction mapping theorem applies, proving d^o_t(z) → d^o(z) for large T.

Inference

Estimating a model when all heterogeneity is observed

 Let v_{jt}(z_t) denote the flow payoff of any action j ∈ {1,..., J} plus the expected future utility of behaving optimally from period t + 1 on:

$$v_{jt}(z_t) \equiv u_{jt}(z_t) + \beta \sum_{z \in Z} V_{t+1}(z) f_{jt}(z|z_t)$$

• By definition:

$$d_{jt}^{o}(z_{t}) \equiv I\left\{v_{jt}(z_{t}) \geq v_{kt}(z_{t}) \forall k\right\}$$

- Suppose we observe the states z_{nt} and decisions $d_{nt} \equiv (d_{n1t}, \ldots, d_{nJt})$ of individuals $n \in \{1, \ldots, N\}$ over time periods $t \in \{1, \ldots, T\}$.
- Could we use such data to infer the primitives of the model:
 - A consistent estimator of $f_{jt}(z_{t+1}|z_t)$ can be obtained from the proportion of observations in the (t, j, z_t) cell transitioning to z_{t+1} .
 - **2** There are $(J-1)\sum_{n=1}^{N} I\{z_{nt} = z_t\}$ inequalities relating the pairs of mappings $v_{jt}(z_t)$ and $v_{kt}(z_t)$ for each observation on d_{nt} at (t, z_t) .
 - **3** Can we recursively derive the values of $u_{jt}(z_t)$ from the $v_{jt}(z_t)$ values?

Why unobserved heterogeneity is introduced into data analysis

- Note that if two people in the data set with the same (t, z_t) made different decisions, say j and k, then $v_{jt}(z_t) = v_{kt}(z_t)$. This raises two potential problems for modeling data this way:
 - In a large data set it is easy to imagine that for every choice j ∈ {1,..., J} and every (t, z_t) at least one sampled person n sets d_{njt} = 1. If so, we would conclude that the population was indifferent between all the choices, and hence the model would have no empirical content because no behavior could be ruled out.
 - This approach does not make use of the information that some choices are more likely than others; that is the proportions of the sample taking different choices at (t, z_t) might vary, some choices being observed often, others perhaps very infrequently.
- For these two reasons, treating all heterogeneity as observed, and trying to predict the decisions of individuals, is not a very promising approach to analyzing data.

Unobserved heterogeneity

- A more modest objective is to predict the probability distribution of choices margined over factors that individuals observe, but data analysts do not.
- In this respect we seek to predict the behavior of a population, not each individual, essentially obliterating that difference between macroeconomics and microeconomics.
- We now assume the states can be partitioned into those which are observed, x_t , and those that are not, ϵ_t .
- Thus $z_t \equiv (x_t, \epsilon_t)$.
- Suppose the data consist of N independent and identically distributed draws from the string of random variables (X₁, D₁,..., X_T, D_T).
- The n^{th} observation is given by $\left\{x_1^{(n)}, d_1^{(n)}, \dots, x_T^{(n)}, d_T^{(n)}\right\}$ for $n \in \{1, \dots, N\}$.

Denote the mixed probability (density) of the pair (x_{t+1}, e_{t+1}), conditional on (x_t, e_t) and the optimal action is j, as:

$$\mathcal{H}_{jt}(x_{t+1},\epsilon_{t+1}|x_t,\epsilon_t) \equiv d_{jt}^o(x_t,\epsilon_t) f_{jt}(x_{t+1},\epsilon_{t+1}|x_t,\epsilon_t)$$

• The probability of $\{d_1, x_2, \dots, d_{T-1}, x_T, d_T\}$ given x_1 is:

$$\Pr\left\{d_{1}, x_{2}, \dots, d_{T-1}, x_{T}, d_{T} | x_{1}\right\} = \int_{\epsilon_{T}} \cdots \int_{\epsilon_{1}} \left[\begin{array}{c} g\left(\epsilon_{1} | x_{1}\right) \sum_{j=1}^{J} d_{jT} d_{jT}^{o}\left(x_{T}, \epsilon_{T}\right) \times \\ \prod_{t=1}^{T-1} \sum_{j=1}^{J} d_{jt} H_{jt}\left(x_{t+1}, \epsilon_{t+1} | x_{t}, \epsilon_{t}\right) \end{array} \right] d\epsilon_{1} \dots d\epsilon_{T}$$

where $g(\epsilon_1 | x_1)$ is the density of ϵ_1 conditional on x_1 .

- Let $\theta \in \Theta$ uniquely index a specification of $u_{jt}(z_t)$, $f_{jt}(z_{t+1}|z_t)$ and β under consideration.
- Conditional on $x_1^{(n)}$ suppose $\left\{d_1^{(n)}, x_2^{(n)}, \ldots, d_T^{(n)}\right\}_{n=1}^N$ was generated by $\theta_0 \in \Theta$.
- Define $\epsilon \equiv (\epsilon_1, \dots, \epsilon_T)$. The maximum likelihood (ML) estimator, θ_{ML} , selects $\theta \in \Theta$ to maximize the joint probability of the observed occurrences conditional on the initial conditions:

$$\theta_{ML} \equiv \operatorname*{arg\,max}_{\theta \in \Theta} \left\{ N^{-1} \sum_{n=1}^{N} \log \left(\Pr\left\{ d_1^{(n)}, x_2^{(n)}, \dots, x_T^{(n)}, d_T^{(n)} \left| x_1^{(n)}; \theta \right. \right\} \right) \right\}$$

Inference

Identification and the properties of the ML estimator

• This model is point identified if and only if (iff) θ_0 is the unique solution when $\theta \in \Theta$ is chosen to maximize:

$$\int_{x_1^{(n)}} \log \left(\Pr\left\{ d_1^{(n)}, x_2^{(n)}, \dots, x_T^{(n)}, d_T^{(n)} \, \middle| \, x_1^{(n)}; \theta \right\} \right) \, dF\left(x_1^{(n)}\right)$$

- If the model is point identified, θ_{ML} is \sqrt{N} consistent, asymptotically normal, and asymptotically efficient:
 - a model is *point identified* if no other model in the Θ set of models has the same *data generating process*.
 - 2 an estimator of an identified model is *consistent* if it converges to θ_0 in some probabilistic sense as N increases without bound.
 - (a) the rate of convergence, 1/2 in this case, is the greatest α leaving the limit of $N^{\alpha} (\theta_{ML} \theta_0)$ bounded in some probabilistic sense.
 - asymptotic normality means the *limiting distribution* (again as N increases without bound), of $\sqrt{N} (\theta_{ML} \theta_0)$ is normal.
 - asymptotic efficiency refers to the lowest asymptotic variance of all consistent estimators with the same rate of convergence.

Criteria for Evaluating Estimators

Three criteria for assessing an estimator

- Three criteria for evaluating an estimator of a point-identified model are:
 - Large sample properties:
 - Does the estimator converge to the identified set?
 - If so, what is the rate of convergence?
 - What is the asymptotic distribution of the estimator?
 - 2 Finite sample properties:
 - At what sample size do the finite sample properties accurately reflect the asymptotic distribution?
 - For a given sample size, what is the standard deviation and mean squared error of the estimator ?
 - Implementation:
 - Is the estimator defined by an algorithm or only a set of conditions to be satisfied?
 - Are numerical approximations involved?
 - Does the estimator require tuning parameters or instruments?

Large Sample or Asymptotic Properties

In what sense does an estimator converge, and what does it converge to?

- There are several types of convergence, such as: almost sure, in mean square, and in probability.
- Given a type of convergence, we ask:
 - Does the estimator converge to a set that includes the identified set? In other words is the estimator tight?
 - Is the set of parameters to which the estimator converges included in the identified set? In other words is the estimator sharp?
- If both conditions are satisfied, then we say the estimator is consistent.
- For example if the identified set is a singleton, that is the model is pointwise identified, then an estimator is consistent if it converges to that singleton.
- Note that if the model is not point identified, we would not expect an extremum estimator (such as a conventionally defined ML) to converge.

- The other two criteria are extensively analyzed in econometric theory, and can typically be applied to dynamic discrete choice models in a straightforward way.
- For example, suppose the parameter space is Θ, the data is generated by θ₀ ∈ Θ, the model in point identified, and the estimator, denoted by θ_N is consistent with:

$$\theta_N \xrightarrow{p} \theta_0$$

• The rate of convergence is defined by N^{α} where:

$$\alpha = \arg\sup_{a} \left[N^{a} \left(\theta_{N} - \theta_{0} \right) \right] \underset{p}{\longrightarrow} 0$$

• Structural estimates of dynamic discrete choice models are typically \sqrt{N} consistent.

Large Sample or Asymptotic Properties

The asymptotic distribution

- Suppose θ_N converges in probability to θ_0 at rate α .
- Let ξ be drawn from the limiting distribution of $N^{\alpha} (\theta_N \theta_0)$:

$$N^{\alpha}\left(\theta_{N}-\theta_{0}
ight)\xrightarrow[d]{d}\xi$$

- Structural estimates of dynamic discrete choice models are typically asymptotically normal.
- An estimator is asymptotically efficient if ξ is $\mathcal{N}\left(0, \mathcal{I}\left(\theta_{0}\right)^{-1}\right)$ where:

$$\mathcal{I}(\theta) \equiv E\left[\frac{\partial I(d, x | x_{1}; \theta)}{\partial \theta} \frac{\partial I(d, x | x_{1}; \theta)'}{\partial \theta}\right] = -E\left[\frac{\partial^{2} I(d, x | x_{1}; \theta)}{\partial \theta \partial \theta'}\right]$$

and the likelihood is based on the sequence (d, x) conditional on the state at date one, x_1 .

• The ML estimator for dynamic discrete choice models typically attain $\mathcal{I}(\theta_0)^{-1}$ the Cramer-Rao lower bound.

- Ideally an estimator is defined by an algorithm that depends on the data for each sample size *N*. In that case the estimator:
 - I can be implemented mechanically, so is easy to explain;
 - is easy to replicate on the same and on different data sets, a virtue in scientific enquiry.
- Cell estimators and hence unrestricted ML estimators satisfy this definition.
- An OLS estimator also satisfies the first definition because algorithms exist to invert matrices exactly, within a finite number of steps.
- Similarly Gaussian methods, successively substituting out parameters, solve linear systems quickly within a finite number of steps.

Is the estimator defined by a set of conditions it must satisfy?

- A weaker, more inclusive definition is that an estimator solves a set of conditions jointly satisfied by the parameter values and the data.
- Since the algorithm used to implement the estimator is not defined, such estimators are almost invariably, less transparent, and therefore harder to replicate with data.
- Extremum estimators for nonlinear models defined this way include:
 - nonlinear least squares;
 - full solution estimators to dynamic discrete choice models;
 - CCP estimators in which G or β is estimated.
- It is useful to know whether a unique solution exists. For example:
 - Is the minimization (maximization) problem strictly convex (concave)?
- If not, can all the parameters, bar one or two, be solved in terms of the one or two remaining parameters?
 - In the first case, the concentrated objective function can be plotted.
 - In the second equi-value contours can be plotted.

- Because ML estimation of dynamic discrete choice models is relatively imposing in terms of programming demands and computational time, researchers economize on both by using numerical approximations:
 - approximating distant horizons with zero;
 - approximating smoothed integrals with rectangles and quadrilaterals;
 - Iinearizing the value function;
 - interpolating the state space to obtain estimates of continuation values;
 - approximating $E[\max{x, y}]$ with $\max{E[x], E[y]}$;
 - reducing the impact of the state space by treating the continuation value as a sufficient statistic for the state space;
 - more generally only allowing the individuals to condition on a smaller set of values than there are state variables.
- These approximation errors open a gap between the defined estimator and its numerical counterpart.

< ロ > < 同 > < 回 > < 回 > < 回