

**SUPPLEMENTARY APPENDIX TO  
“WINNING BY DEFAULT: WHY IS THERE SO LITTLE  
COMPETITION IN GOVERNMENT PROCUREMENT?”**

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In this document, we provide the proof of the lemmas and the theorem in Section 5 on the nonparametric identification of the model, the closed-form characterization of the equilibrium contracts and competition of the estimated model, and the details of the simulated GMM estimator.

1. PROOF OF LEMMA 5.1

**Lemma 5.1 (i)** If **A3** holds then  $\partial |q(l, \pi)| / \partial \pi > 0$ .

*Proof.* Theorem 4.1 characterizes the unique optimal menu of two contracts, defined by the price of the fixed contract:

$$\underline{p}_n = \alpha + \frac{\pi(1-\pi)^{n-1}}{1-(1-\pi)^n} \left( \beta - \int \psi[q(s)] [1-l(s)] \bar{f}(s) ds \right), \quad (1)$$

and the variable contract:

$$\bar{p} = \alpha + \beta - \int \psi[q(s)] \bar{f}(s) ds, \quad (2)$$

$$q(s) = \begin{cases} h \left( \frac{1-\min\{\pi, \tilde{\pi}\}}{1-\min\{\pi, \tilde{\pi}\}l(s)} \right) & \text{if } l(s) \leq \tilde{l}(\min\{\pi, \tilde{\pi}\}), \\ M & \text{if } l(s) > \tilde{l}(\min\{\pi, \tilde{\pi}\}). \end{cases} \quad (3)$$

By **A3**  $q(l, \pi)$  satisfies:

$$\psi' [q(l, \pi)] [1 - \pi l] = 1 - \pi. \quad (4)$$

Totally differentiating (4) with respect to  $\pi$  and making  $\partial q(l, \pi) / \partial \pi$  the subject of the resulting equation:

$$\frac{\partial q(l, \pi)}{\partial \pi} = \frac{l - 1}{\psi'' [q(l, \pi)] (1 - \pi l)^2}. \quad (5)$$

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By assumption  $\psi''(q) < 0$ , implying  $\partial q(l, \pi) / \partial \pi \gtrless 0$  when  $l \gtrless 1$ . From (4) it follows that  $q(l, \pi) \gtrless 0$  when  $l \gtrless 1$ . Combining both sets of inequalities  $\partial q(l, \pi) / \partial \pi \gtrless 0$  when  $q(l, \pi) \gtrless 0$ , as claimed.  $\square$

**Lemma 5.1 (ii)** If **A3** and **A5** hold then  $\partial \bar{p}(\pi) / \partial \pi < 0$ .

*Proof.* Totally differentiating the base price for variable contracts, defined in (2), with respect to  $\pi$ , yields:

$$\bar{p}'(\pi) = \alpha'(\pi) + \beta'(\pi) - \int \psi'(q[l(s), \pi]) \frac{\partial q[l(s), \pi]}{\partial \pi} \bar{f}(s) ds.$$

Define:

$$m(l, \pi) \equiv \psi'[q(l, \pi)] \frac{\partial q(l, \pi)}{\partial \pi}.$$

From (7) the first order condition for an interior solution can be rewritten as:

$$\psi'(q) = (1 - \pi) / (1 - \pi l), \text{ or } q(l, \pi) = h[(1 - \pi) / (1 - \pi l)].$$

Therefore:

$$m(l, \pi) = h' \left( \frac{1 - \pi}{1 - \pi l} \right) \frac{(1 - \pi)(1 - l)}{(1 - \pi l)^3}.$$

Note  $m(l, \pi) \gtrless 0$  for  $l \gtrless 1$  because  $h'(x) < 0$  and by **A3**,  $\pi l < 1$ . By definition  $l(s) > 0$  for all  $s \in S$ . Therefore:

$$\begin{aligned} \bar{p}'(\pi) &= \alpha'(\pi) + \beta'(\pi) - \int m[l(s), \pi] \bar{f}(s) ds \\ &< \alpha'(\pi) + \beta'(\pi) - \int_{\{s:l(s)<1\}} m[l(s), \pi] \bar{f}(s) ds \\ &\leq \alpha'(\pi) + \beta'(\pi) - \sup_{l \in (0,1)} |m(l, \pi)| \int_{\{s:l(s)<1\}} \bar{f}(s) ds. \end{aligned}$$

Thus  $|m(l, \pi)|$  is bounded for  $l \in (0, 1)$  by:

$$\begin{aligned} |m(l, \pi)| &\leq \left| h' \left( \frac{1 - \pi}{1 - \pi l} \right) \right| \left| \frac{(1 - \pi)(1 - l)}{(1 - \pi l)^3} \right| \\ &\leq \sup_{l \in (0,1)} \left| h' \left( \frac{1 - \pi}{1 - \pi l} \right) \right| \sup_{l \in (0,1)} (1 - l) \sup_{l \in (0,1)} \left| \frac{(1 - \pi)}{(1 - \pi l)^3} \right| \\ &\leq (1 - \pi)^{-2} \sup_{l \in (0,1)} \left| h' \left( \frac{1 - \pi}{1 - \pi l} \right) \right|. \end{aligned}$$

Therefore:

$$\bar{p}'(\pi) < \alpha'(\pi) + \beta'(\pi) - (1 - \pi)^{-2} \sup_{l \in (0,1)} \left| h' \left( \frac{1 - \pi}{1 - \pi l} \right) \right| \int_{\{s:l(s)<1\}} \bar{f}(s) ds < 0.$$

where the second inequality holds by **A5**.  $\square$

**Lemma 5.1 (iii)** If **A3** and **A4** hold then  $\partial \underline{p}_n(\pi) / \partial \pi < 0$  for all  $n \in \{1, 2, \dots\}$ .

*Proof.* Rewriting (1) to make the dependence of  $\underline{p}_n$  on  $\pi$  explicit:

$$\begin{aligned} \underline{p}_n(\pi) &= \alpha(\pi) + \frac{\pi(1-\pi)^{n-1}}{1-(1-\pi)^n} \left[ \beta(\pi) - \int \psi(q[l(s), \pi]) [1-l(s)] \bar{f}(s) \right] \\ &\equiv \alpha(\pi) + \Psi_{0,n}(\pi) [\beta(\pi) - \Psi_1(\pi)]. \end{aligned}$$

and hence:

$$\underline{p}'_n(\pi) = \alpha'(\pi) + \Psi'_{0,n}(\pi) [\beta(\pi) - \Psi_1(\pi)] + \Psi_{0,n}(\pi) [\beta'(\pi) - \Psi'_1(\pi)]. \quad (6)$$

By **A4**  $\alpha'(\pi) \leq 0$ . Also:

$$\frac{\partial}{\partial \pi} \ln [\Psi_{0,n}(\pi)] = \frac{1 - n\pi - (1-\pi)^n}{\pi(1-\pi)[1-(1-\pi)^n]}.$$

The derivative is zero at  $n = 1$  and  $-\pi^2$  at  $n = 2$ . Now suppose it is negative for all  $n \in \{2, \dots, n_0\}$ . For  $n_0 + 1$  the denominator is positive and the numerator is:

$$1 - (n_0 + 1)\pi - (1-\pi)(1-\pi)^{n_0} < \pi(1-\pi)^{n_0} - \pi < 0.$$

The first inequality follows from an induction hypothesis, and the second one from the inequalities  $0 < \pi < 1$ . Therefore  $\Psi'_{0,n}(\pi) \leq 0$  for all  $(\pi, n)$ . By **A3** the participation constraint is satisfied with an inequality implying  $\beta(\pi) > \Psi_1(\pi)$ , and hence  $\Psi'_{0,n}(\pi) [\beta(\pi) - \Psi_1(\pi)] \leq 0$ . The third expression in (6) is strictly negative because  $\Psi_{0,n}(\pi) > 0$ , by **A4**  $\beta'(\pi) \leq 0$ , and:

$$\begin{aligned} \Psi'_1(\pi) &= \int \psi'(q[l(s), \pi]) \frac{\partial q[l(s), \pi]}{\partial \pi} [1-l(s)] \bar{f}(s) ds \\ &= \int (1-\pi) \frac{\partial q[l(s), \pi]}{\partial \pi} \left[ \frac{1-l(s)}{1-\pi l(s)} \right] \bar{f}(s) ds \\ &= \int \frac{(\pi-1)[1-l(s)]^2}{\psi''(q[l(s), \pi])[1-\pi l(s)]^3} \bar{f}(s) ds > 0. \end{aligned}$$

Appealing to **A3**, the second equality uses (4) to substitute out  $\psi'[q(s, \pi)]$ , the third equality uses (5) to substitute out  $\partial q(l, \pi) / \partial \pi$ , and the inequality follows from  $\psi''(q) < 0$  and the assumption of an interior solution. Therefore  $\underline{p}'_n(\pi) < 0$ .  $\square$

## 2. PROOF OF LEMMA 5.2

**Lemma 5.2**  $f_{\pi|c,n,v}(\pi|c, n, v)$  is identified.

*Proof.* Rewriting (4) to solve for  $\pi$ , we obtain:

$$\pi = [1 - \psi'(q)] / [1 - \psi'(q)l]. \quad (7)$$

Since  $\psi(q)$  is identified, the  $\pi$  corresponding to each variable contract  $(\bar{p}, q, l)$  is identified from (7). Hence  $f_{\pi|c,n,v}(\pi|c, n, 1)$  is identified. Define the odds ratio related to contract types conditional on  $(c, n)$  as:

$$\varphi_{c,n} \equiv \Pr(v = 1|c, n) / \Pr(v = 0|c, n).$$

and note that  $\varphi_{c,n}$  is identified. The joint probability that the contract type is fixed and  $\pi \leq \pi^*$  can be expressed as:

$$\begin{aligned} \Pr\{\pi \leq \pi^*, v = 0|c, n\} &= F_{\pi|c,n,v}(\pi^*|c, n, 0) \Pr(v = 0|c, n) \\ &= \int_{\pi=\underline{\pi}}^{\pi^*} f_{\pi|c,n}(\pi|c, n) [1 - (1 - \pi)^n] d\pi. \end{aligned} \quad (8)$$

Taking the derivative with respect to  $\pi^*$  yields:

$$f_{\pi|c,n,v}(\pi^*|c, n, 0) \Pr(v = 0|c, n) = f_{\pi|c,n}(\pi^*|c, n) [1 - (1 - \pi^*)^n]. \quad (9)$$

Similarly:

$$\begin{aligned} \Pr\{\pi \leq \pi^*, v = 1|c, n\} &= F_{\pi|c,n,v}(\pi^*|c, n, 1) \Pr(v = 1|c, n) \\ &= \int_{\pi=\underline{\pi}}^{\pi^*} f_{\pi|c,n}(\pi|c, n) (1 - \pi)^n d\pi, \end{aligned}$$

and taking the derivative with respect to  $\pi^*$  yields:

$$f_{\pi|c,n,v}(\pi^*|c, n, 1) \Pr(v = 1|c, n) = f_{\pi|c,n}(\pi^*|c, n) (1 - \pi^*)^n. \quad (10)$$

Rearranging the quotient of (9) and (10) to make  $f_{\pi|c,n,v}(\pi^*|c, n, v = 0)$  the subject of the resulting equation, we obtain:

$$f_{\pi|c,n,v}(\pi|c, n, 0) = \varphi_{c,n} \frac{[1 - (1 - \pi)^n]}{(1 - \pi)^n} f_{\pi|c,n,v}(\pi|c, n, 1),$$

after dropping the superscripts asterisks. Because  $\varphi_{c,n}$  and  $f_{\pi|c,n,v}(\pi|c, n, 1)$  are identified, so is  $f_{\pi|c,n,v}(\pi|c, n, 0)$ .  $\square$

## 3. PROOF OF THEOREM 5.1

**Theorem 5.1**  $\psi(q)$ ,  $\alpha(\pi)$  and  $\beta(\pi)$  are identified, and for  $n \in \{2, 3, \dots\}$ :

$$\begin{aligned}\alpha(\pi) &= \frac{1 - (1 - \pi)^n}{1 - (1 - \pi)^{n-1}} \underline{p}_n^*(\pi, c) - \frac{\pi(1 - \pi)^{n-1}}{1 - (1 - \pi)^{n-1}} \underline{p}_1^*(\pi, c), \\ \beta(\pi) &= \bar{p}(\pi) + \int \psi \left( h \left[ \frac{1 - \pi}{1 - \pi l(t)} \right] \right) \bar{f}(t) dt - \alpha(\pi).\end{aligned}\quad (11)$$

*Proof.* To prove the first equation in (11) set  $n = 1$  in (1) to obtain:

$$\underline{p}_1 - \alpha = \beta - \int \psi[q(s)] [1 - l(s)] \bar{f}(s) ds.$$

Substitute the right hand side of this expression back into (1) yielding:

$$\underline{p}_n = \alpha + \frac{\pi(1 - \pi)^{n-1}}{1 - (1 - \pi)^n} (\underline{p}_1 - \alpha).$$

Make  $\alpha$  the subject of the equation, and explicitly recognize the dependence of  $\underline{p}_n$ ,  $\underline{p}_1$  and  $\alpha$  on  $\pi$ . To prove the second equation in (11), rearrange (7) using the definition of  $h(\cdot)$  to obtain  $q$  as a function of  $s$ , specifically  $q(s) = h((1 - \pi)/[1 - \pi l(s)])$ , substitute the expression into (2), make  $\beta$  the subject, and the dependence of  $\alpha$ ,  $\beta$  and  $\bar{p}$  on  $\pi$  explicit.

The proof that  $\psi(q)$  is identified is given in the text. Proving  $\alpha(\pi)$  and  $\beta(\pi)$  are identified follows from noting that their representations given in (11) are identified by Lemma 5.2 and the arguments in the surrounding text.  $\square$

## 4. OPTIMAL COMPETITION UNDER THE PARAMETRIC SPECIFICATION

The specification on  $\kappa(\pi, n)$  is

$$\kappa(n, \pi) = (\kappa_1 + \kappa_2\pi)(n - 1) + (\kappa_3 + \kappa_4\pi)(n - 1)^2.$$

Given this specification, the expected total cost of competed procurement with effort  $\lambda$ , denoted by  $U(\pi, \lambda)$ , is:

$$\begin{aligned}U(\pi, \lambda) &= \alpha(\pi) + \exp^{-\lambda\pi} [\beta(\pi) + \Gamma(\pi)] + \mathbb{E}[\kappa(\pi, j + 1)|\lambda] + \eta \\ &= \alpha(\pi) + \exp^{-\lambda\pi} [\beta(\pi) + \Gamma(\pi)] + \tilde{\kappa}_1(\pi)\lambda + \tilde{\kappa}_2(\pi)\lambda(1 + \lambda) + \eta,\end{aligned}$$

where  $\tilde{\kappa}_1(\pi) \equiv \kappa_1 + \kappa_2\pi$  and  $\tilde{\kappa}_2(\pi) \equiv \kappa_3 + \kappa_4\pi$ .

Taking the first order condition:

$$\pi \exp^{-\lambda\pi} [\beta(\pi) + \Gamma(\pi)] = \tilde{\kappa}_1(\pi) + \tilde{\kappa}_2(\pi)(1 + 2\lambda).\quad (12)$$

Because the left hand side is decreasing in  $\lambda$  while the right hand side is increasing in  $\lambda$ , there exists a unique solution to the above equation for any given  $\pi$ , denoted by  $\tilde{\lambda}(\pi)$ . Because  $\lambda \geq 0$ ,  $\lambda^*(\pi) = \max\{\tilde{\lambda}(\pi), 0\}$ . In our estimation, we numerically solve for  $\lambda^*(\pi)$  for each  $\pi$ .

Given  $\lambda^*(\pi)$ , it is optimal for the buyer to hold a competitive solicitation if and only if

$$U[\pi, \lambda^*(\pi)] \leq U_0(\pi).$$

The above inequality can be rewritten as:

$$\eta \leq (1 - e^{-\lambda^*(\pi)\pi}[\beta(\pi) + \Gamma(\pi)] - \tilde{\kappa}_1\lambda^*(\pi) - \tilde{\kappa}_2\lambda^*(\pi)[1 + \lambda^*(\pi)]). \quad (13)$$

Equations (1), (2), (3), (12), and (13) characterize the equilibrium contracts and competition.

## 5. SIMULATED GMM ESTIMATOR

Let us denote the vector of the parameters of the model by  $\theta$ . Our estimator minimizes a weighted sum of squared distances:

$$g_n(\theta)'Wg_n(\theta), \text{ with } g_n(\theta) = \frac{1}{n} \sum_{t=1}^n g(w_t; \theta),$$

where  $W$  is a symmetric positive-definite weighting matrix. The  $g(w_t; \theta)$  vector is associated with 40 moment conditions: (i) 17 moment conditions on competition, contract type, and contract price for all projects, (ii) the same 17 moment conditions for non-military projects, and (iii) 6 moment conditions on the distribution of  $s$ , or the standardized delay.

The 17 moment conditions consist of  $\Pr(c_i = 0)$ ,  $\Pr(n_i = 1|c_i = 1)$ ,  $\Pr(n_i \leq 2|c_i = 1)$ ,  $\Pr(n_i \leq 5|c_i = 1)$ ,  $\Pr(v_i = 0|c_i = 0)$ ,  $\Pr(v_i = 0|c_i = 1)$ ,  $\Pr(v_i = 0|c_i = 1, n_i = 1)$ ,  $\Pr(v_i = 0|c_i = 1, n_i \leq 2)$ ,  $\Pr(v_i = 0|c_i = 1, n_i \leq 5)$ ,  $\mathbb{E}[n_i]$ ,  $\mathbb{E}[p_i|c_i = c, v_i = v]$  for  $(c, v) \in \{0, 1\} \times \{0, 1\}$  where  $p_i \equiv \underline{p}_i(1 - v_i) + (\bar{p}_i + q_i)v_i$ ,  $\mathbb{E}[p_i|c_i = 1, n_i = 1]$ ,  $\mathbb{E}[p_i|c_i = 1, n_i \leq 2]$ , and  $\mathbb{E}[p_i|c_i = 1, n_i \leq 5]$ . The 6 moment conditions on the distribution of the standardized delay are  $\Pr(s_i = 0, v_i = 0)$ ,  $\Pr(s_i = 0, v_i = 1)$ ,  $\Pr[s_i(1 - v_i)]$ ,  $\Pr[s_i^2(1 - v_i)]$ ,  $\Pr[s_i v_i]$ , and  $\Pr[s_i^2 v_i]$ . Note that the moments as a function of  $\theta$  are calculated using simulation. In our estimation, the simulation size is 5,000.

We use a two-step procedure to obtain the optimally weighted simulated GMM estimator. We start with a positive definite weighting matrix and obtain a first-step

estimator, denoted by  $\tilde{\theta}_n$ . The asymptotic variance of  $\sqrt{n}g_n(\theta_0)$ ,  $S$ , is estimated by:

$$\hat{S} \equiv \frac{1}{n} \sum_t g(w_t, \tilde{\theta}_n) g(w_t; \tilde{\theta}_n)'$$

Then we re-estimate the parameters using the optimal weighting matrix  $\hat{S}^{-1}$  to obtain the optimally weighted simulated GMM estimator, which we denote by  $\hat{\theta}_n$ .

Under standard regularity conditions, this estimator is asymptotically normally distributed, and a consistent estimator of the asymptotic variance of  $\sqrt{n}(\hat{\theta}_n - \theta_0)$  is:

$$\left( \frac{\partial g_n(\hat{\theta}_n)}{\partial \theta'}' \hat{S}^{-1} \frac{\partial g_n(\hat{\theta}_n)}{\partial \theta'} \right)^{-1}.$$

Since the moments are calculated by simulation, we use a numerical derivative of  $g_n(\theta)$  to estimate the asymptotic variance of the estimator.